

8.14 The CLT tells us that $Z = (\sum_i X_i - np) / \sqrt{np(1-p)}$ is approximately $n(0, 1)$. For a test that rejects H_0 when $\sum_i X_i > c$, we need to find c and n to satisfy

$$P\left(Z > \frac{c - n(.49)}{\sqrt{n(.49)(.51)}}\right) = .01 \quad \text{and} \quad P\left(Z > \frac{c - n(.51)}{\sqrt{n(.51)(.49)}}\right) = .99.$$

We thus want

$$\frac{c - n(.49)}{\sqrt{n(.49)(.51)}} = 2.33 \quad \text{and} \quad \frac{c - n(.51)}{\sqrt{n(.51)(.49)}} = -2.33.$$

Solving these equations gives $n = 13,567$ and $c = 6,783.5$.

8.15 From the Neyman-Pearson lemma the UMP test rejects H_0 if

$$\frac{f(x | \sigma_1)}{f(x | \sigma_0)} = \frac{(2\pi\sigma_1^2)^{-n/2} e^{-\sum_i x_i^2 / (2\sigma_1^2)}}{(2\pi\sigma_0^2)^{-n/2} e^{-\sum_i x_i^2 / (2\sigma_0^2)}} = \left(\frac{\sigma_0}{\sigma_1}\right)^n \exp\left\{\frac{1}{2} \sum_i x_i^2 \left(\frac{1}{\sigma_0^2} - \frac{1}{\sigma_1^2}\right)\right\} > k$$

for some $k \geq 0$. After some algebra, this is equivalent to rejecting if

$$\sum_i x_i^2 > \frac{2 \log(k(\sigma_1/\sigma_0)^n)}{\left(\frac{1}{\sigma_0^2} - \frac{1}{\sigma_1^2}\right)} = c \quad \left(\text{because } \frac{1}{\sigma_0^2} - \frac{1}{\sigma_1^2} > 0\right).$$

This is the UMP test of size α , where $\alpha = P_{\sigma_0}(\sum_i X_i^2 > c)$. To determine c to obtain a specified α , use the fact that $\sum_i X_i^2 / \sigma_0^2 \sim \chi_n^2$. Thus

$$\alpha = P_{\sigma_0}\left(\sum_i X_i^2 / \sigma_0^2 > c / \sigma_0^2\right) = P(\chi_n^2 > c / \sigma_0^2),$$

so we must have $c / \sigma_0^2 = \chi_{n,\alpha}^2$, which means $c = \sigma_0^2 \chi_{n,\alpha}^2$.

8.16 a.

$$\begin{aligned} \text{Size} &= P(\text{reject } H_0 | H_0 \text{ is true}) = 1 \Rightarrow \text{Type I error} = 1. \\ \text{Power} &= P(\text{reject } H_0 | H_A \text{ is true}) = 1 \Rightarrow \text{Type II error} = 0. \end{aligned}$$

b.

$$\begin{aligned} \text{Size} &= P(\text{reject } H_0 | H_0 \text{ is true}) = 0 \Rightarrow \text{Type I error} = 0. \\ \text{Power} &= P(\text{reject } H_0 | H_A \text{ is true}) = 0 \Rightarrow \text{Type II error} = 1. \end{aligned}$$

8.19 The pdf of Y is

$$f(y|\theta) = \frac{1}{\theta} y^{(1/\theta)-1} e^{-y^{1/\theta}}, \quad y > 0.$$

By the Neyman-Pearson Lemma, the UMP test will reject if

$$\frac{1}{2} y^{-1/2} e^{y-y^{1/2}} = \frac{f(y|2)}{f(y|1)} > k.$$

To see the form of this rejection region, we compute

$$\frac{d}{dy} \left(\frac{1}{2} y^{-1/2} e^{y-y^{1/2}} \right) = \frac{1}{2} y^{-3/2} e^{y-y^{1/2}} \left(y - \frac{y^{1/2}}{2} - \frac{1}{2} \right)$$

which is negative for $y < 1$ and positive for $y > 1$. Thus $f(y|2)/f(y|1)$ is decreasing for $y \leq 1$ and increasing for $y \geq 1$. Hence, rejecting for $f(y|2)/f(y|1) > k$ is equivalent to rejecting for $y \leq c_0$ or $y \geq c_1$. To obtain a size α test, the constants c_0 and c_1 must satisfy

$$\alpha = P(Y \leq c_0 | \theta = 1) + P(Y \geq c_1 | \theta = 1) = 1 - e^{-c_0} + e^{-c_1} \quad \text{and} \quad \frac{f(c_0|2)}{f(c_0|1)} = \frac{f(c_1|2)}{f(c_1|1)}.$$

Solving these two equations numerically, for $\alpha = .10$, yields $c_0 = .076546$ and $c_1 = 3.637798$. The Type II error probability is

$$P(c_0 < Y < c_1 | \theta = 2) = \int_{c_0}^{c_1} \frac{1}{2} y^{-1/2} e^{-y^{1/2}} dy = -e^{-y^{1/2}} \Big|_{c_0}^{c_1} = .609824.$$

8.49 a. The p-value is

$$\begin{aligned} P\left\{\left(\begin{array}{c} 7 \text{ or more successes} \\ \text{out of 10 Bernoulli trials} \end{array}\right) \middle| \theta = \frac{1}{2}\right\} \\ = \binom{10}{7} \left(\frac{1}{2}\right)^7 \left(\frac{1}{2}\right)^3 + \binom{10}{8} \left(\frac{1}{2}\right)^8 \left(\frac{1}{2}\right)^2 + \binom{10}{9} \left(\frac{1}{2}\right)^9 \left(\frac{1}{2}\right)^1 + \binom{10}{10} \left(\frac{1}{2}\right)^{10} \left(\frac{1}{2}\right)^0 \\ = .171875. \end{aligned}$$

b.

$$\begin{aligned} \text{P-value} &= P\{X \geq 3 \mid \lambda = 1\} = 1 - P(X < 3 \mid \lambda = 1) \\ &= 1 - \left[\frac{e^{-1}1^2}{2!} + \frac{e^{-1}1^1}{1!} + \frac{e^{-1}1^0}{0!} \right] \approx .0803. \end{aligned}$$

c.

$$\begin{aligned} \text{P-value} &= P\left\{\sum_i X_i \geq 9 \mid 3\lambda = 3\right\} = 1 - P(Y < 9 \mid 3\lambda = 3) \\ &= 1 - e^{-3} \left[\frac{3^8}{8!} + \frac{3^7}{7!} + \frac{3^6}{6!} + \frac{3^5}{5!} + \cdots + \frac{3^1}{1!} + \frac{3^0}{0!} \right] \approx .0038, \end{aligned}$$

where $Y = \sum_{i=1}^3 X_i \sim \text{Poisson}(3\lambda)$.

10.32 a. First calculate the MLEs under $p_1 = p_2 = p$. We have

$$L(p|x) = p^{x_1} p^{x_2} p^{x_3} \cdots p^{x_{n-1}} \left(1 - 2p - \sum_{i=3}^{n-1} p_i\right)^{m - x_1 - x_2 - \cdots - x_{n-1}}$$

Taking logs and differentiating yield the following equations for the MLEs:

$$\begin{aligned} \frac{\partial \log L}{\partial p} &= \frac{x_1 + x_2}{p} - \frac{2(m - \sum_{i=1}^{n-1} x_i)}{1 - 2p - \sum_{i=3}^{n-1} p_i} = 0 \\ \frac{\partial \log L}{\partial p_i} &= \frac{x_i}{p_i} - \frac{x_n}{1 - 2p - \sum_{i=3}^{n-1} p_i} = 0, \quad i = 3, \dots, n-1, \end{aligned}$$

with solutions $\hat{p} = \frac{x_1 + x_2}{2m}$, $\hat{p}_i = \frac{x_i}{m}$, $i = 3, \dots, n-1$, and $\hat{p}_n = (m - \sum_{i=1}^{n-1} x_i) / m$. Except for the first and second cells, we have expected = observed, since both are equal to x_i . For the first two terms, expected = $m\hat{p} = (x_1 + x_2)/2$ and we get

$$\sum \frac{(\text{observed} - \text{expected})^2}{\text{expected}} = \frac{(x_1 - \frac{x_1 + x_2}{2})^2}{\frac{x_1 + x_2}{2}} + \frac{(x_2 - \frac{x_1 + x_2}{2})^2}{\frac{x_1 + x_2}{2}} = \frac{(x_1 - x_2)^2}{x_1 + x_2}.$$

b. Now the hypothesis is about conditional probabilities is given by H_0 : $P(\text{change—initial agree}) = P(\text{change—initial disagree})$ or, in terms of the parameters $H_0 : \frac{p_1}{p_1 + p_3} = \frac{p_2}{p_2 + p_4}$. This is the same as $p_1 p_4 = p_2 p_3$, which is not the same as $p_1 = p_2$.